

Data Assimilation: Interacting particle & mean-field formulations

Sebastian Reich

University of Potsdam & SFB 1294 Data Assimilation

September 29, 2023

UQ & DA, UP & SFB 1294

Content



- Part I. Introduction to data assimilation (DA)
 - What is DA?
 - Examples
 - Applications and challenges
- Part II. The ensemble Kalman filter (EnKF)
 - Standard Kalman filter
 - Mean-field EnKF formulation & Monte Carlo implementation
 - Why the EnKF?
 - Why not?
- Part III. Control & DA
 - Coupling of measures
 - Homotopy and data-driven control

<u>Reference</u>: Ensemble Kalman Methods: A Mean Field Perspective; Edoardo Calvello, SR, Andrew M. Stuart, arXiv:2209.11371.



Data assimilation: Two sets of (random) variables (RV) $\{v_n\}$ and $\{\hat{v}_n\}$, $n \ge 0$, which are coupled through the following two alternating steps

(i) **Prediction**:

$$v_n \rightarrow \hat{v}_{n+1}$$

(ii) Analysis:

$$\hat{v}_{n+1} \rightarrow v_{n+1}.$$

Remarks

- $v_n \& \hat{v}_n$ typically take values in \mathbb{R}^d .
- Both continuous and discrete time DA formulations exist.
- The iteration index *n* can represent physical or algorithmic time.



Probabilistic: Sequence of RVs v_n and \hat{v}_n with associated PDFs:

$$m{v}_{n} \sim \mu_{n} \propto e^{-\Psi_{n}}
onumber \ \hat{v}_{n} \sim \hat{\mu}_{n} \propto e^{-\hat{\Psi}_{n}}$$

Work at level of RVs and stochastic processes.



Probabilistic: Sequence of RVs v_n and \hat{v}_n with associated PDFs:

$$m{v}_{n}\sim \mu_{n}\propto e^{-\Psi_{n}}$$

 $\hat{m{v}}_{n}\sim \hat{\mu}_{n}\propto e^{-\hat{\Psi}_{n}}$

Work at level of RVs and stochastic processes.

Optimisation: Sequence of cost functions Φ_n and $\hat{\Phi}_n$ and induced deterministic variables

$$v_n = \arg\min_v \Phi_n(v)$$

 $\hat{v}_n = \arg\min_{\hat{v}} \hat{\Phi}_n(\hat{v})$

Work at level of cost functions (cost-to-go) and optimisation.

SFB 1294 Data Assimilation

Probabilistic: Sequence of RVs v_n and \hat{v}_n with associated PDFs:

$$m{v}_{n}\sim \mu_{n}\propto e^{-\Psi_{n}}$$
 $\hat{m{v}}_{n}\sim \hat{\mu}_{n}\propto e^{-\hat{\Psi}_{n}}$

Work at level of RVs and stochastic processes.

Optimisation: Sequence of cost functions Φ_n and $\hat{\Phi}_n$ and induced deterministic variables

$$v_n = \arg\min_v \Phi_n(v)$$

 $\hat{v}_n = \arg\min_{\hat{v}} \hat{\Phi}_n(\hat{v})$

Work at level of cost functions (cost-to-go) and optimisation.

Remark. In linear Gaussian setting $\Phi_n = \Psi_n$ and $\hat{\Phi}_n = \hat{\Psi}_n$. Otherwise both approaches lead to different results, in general.



A) Markov process: Stochastic process $\{v_n\}$ defined by

 $v_{n+1} = F(v_n, \xi_n)$

with $\{\xi_n\}$ i.i.d. Gaussian and $F(v, \xi)$ a given map. **Task**: Predict v_n given v_0 .



A) Markov process: Stochastic process $\{v_n\}$ defined by

 $v_{n+1} = F(v_n, \xi_n)$

with $\{\xi_n\}$ i.i.d. Gaussian and $F(v, \xi)$ a given map. **Task**: Predict v_n given v_0 .

B) Bayesian inverse problem: Observation model

$$y = h(\hat{v}) + \eta$$

with η Gaussian <u>measurement error</u> and $h(\hat{v})$ a given <u>forward map</u>. **Task**: Infer \hat{v} given observed $y = y^{\dagger}$ and a prior $\hat{\mu}$.



DA formulation: Given a sequence of <u>observations</u> $\{y_n^{\dagger}\}, n \ge 1$:

(i) **Prediction step**: Markov process with $v_n \sim \mu_n$

 $\hat{v}_{n+1}=F(v_n,\xi_n).$



DA formulation: Given a sequence of <u>observations</u> $\{y_n^{\dagger}\}, n \ge 1$:

(i) **Prediction step**: Markov process with $v_n \sim \mu_n$

$$\hat{v}_{n+1}=F(v_n,\xi_n).$$

(ii) Analysis step: Bayesian inference with prior $\hat{v}_{n+1} \sim \hat{\mu}_{n+1}$, negative log-likelihood

$$I_{n+1}(v) = \frac{1}{2} \left(h(v) - y_{n+1}^{\dagger} \right)^{\mathrm{T}} R^{-1} \left(h(v) - y_{n+1}^{\dagger} \right),$$

and posterior

$$\mathbf{v}_{n+1} \sim \mu_{n+1}, \qquad \mu_{n+1}(\mathbf{v}) = \frac{e^{-l_{n+1}(\mathbf{v})}\hat{\mu}_{n+1}(\mathbf{v})}{\hat{\mu}_{n+1}[e^{-l_{n+1}}]}$$



Optimisation: Given a cost function U(v) find the minimiser

$$v^{\dagger} = rg\min_{v \in \mathbb{R}^d} U(v).$$



Optimisation: Given a cost function U(v) find the minimiser

$$v^{\dagger} = rg\min_{v \in \mathbb{R}^d} U(v).$$

Gradient descent:

$$v_{n+1} = v_n - \tau \nabla_v U(v_n)$$

with step-sizes $\tau > 0$.

SFB 1294

Optimisation: Given a cost function U(v) find the minimiser

$$v^{\dagger} = rg\min_{v \in \mathbb{R}^d} U(v).$$

Gradient descent:

$$v_{n+1} = v_n - \tau \nabla_v U(v_n)$$

with step-sizes $\tau > 0$.

Remarks

- Requires derivatives.
- Local minima.
- Not affine invariant (unlike Newton's method).



Probabilistic optimisation: Given a step-size $\tau > 0$ & $v_0 \sim \mu_0$,

(i) **Prediction step**: $v_n \sim \mu_n$

$$\hat{\mathbf{v}}_{n+1} = \mathbf{v}_n + \sqrt{\epsilon \tau} \, \xi_n, \quad \epsilon > \mathbf{0},$$



Probabilistic optimisation: Given a step-size $\tau > 0$ & $v_0 \sim \mu_0$,

(i) **Prediction step**: $v_n \sim \mu_n$

$$\hat{\mathbf{v}}_{n+1} = \mathbf{v}_n + \sqrt{\epsilon \tau} \, \xi_n, \quad \epsilon > \mathbf{0},$$

(i) Analysis step: Prior $\hat{v}_{n+1} \sim \hat{\mu}_{n+1}$, negative log likelihood function l(v) = U(v), and posterior

$$\mathbf{v}_{n+1} \sim \mu_{n+1}, \qquad \mu_{n+1}(\mathbf{v}) = \frac{e^{-\tau U(\mathbf{v})}\hat{\mu}_{n+1}(\mathbf{v})}{\hat{\mu}_{n+1}[e^{-\tau U}]}$$

Remark: Desired

$$\lim_{n\to\infty}v_n\to v^{\dagger}.$$



Alternative affine invariant prediction step:

$$\hat{v}_{n+1} = v_n + \frac{\epsilon \tau}{2} (v_n - m_n^v), \qquad \epsilon > 0$$

with mean

$$m_n^{\mathsf{v}} := \mu_n[\mathsf{v}] = \mathbb{E}_n[\mathsf{v}].$$

Remarks:

(i) mean update

$$\hat{m}_{n+1}^{v}=m_{n}^{v}.$$

(ii) covariance matrix update

$$\hat{C}_{n+1}^{\nu\nu} = \left(1 + \frac{\epsilon\tau}{2}\right)^2 C_n^{\nu\nu}.$$



Denoising diffusion models

Forward / noising process



Figure: Courtesy: scholar.havard.edu



Diffusion modelling: Given samples $x_0^{(j)}$, j = 1, ..., J, from an unknown data distribution p_0 . <u>Noising step</u>: $t \in [0, 1]$,

$$\mathrm{d} x_t = -\frac{1}{2}\beta(t)x_t\mathrm{d} t + \sqrt{\beta(t)}\mathrm{d} w_t, \qquad x_0 = x_0^{(j)} \sim p_0,$$

with transition kernel

$$p_t(x_t|x_0) = N(\sqrt{\alpha_t}x_0, v_t I), \ \alpha_t = \exp\left(-\int_0^t \beta(s) \mathrm{d}s\right), \ v_t = 1 - \alpha_t$$



Diffusion modelling: Given samples $x_0^{(j)}$, j = 1, ..., J, from an unknown data distribution p_0 . <u>Noising step</u>: $t \in [0, 1]$,

$$\mathrm{d} x_t = -\frac{1}{2}\beta(t)x_t\mathrm{d} t + \sqrt{\beta(t)}\mathrm{d} w_t, \qquad x_0 = x_0^{(j)} \sim p_0,$$

with transition kernel

$$p_t(x_t|x_0) = N(\sqrt{\alpha_t}x_0, v_t I), \ \alpha_t = \exp\left(-\int_0^t \beta(s) ds\right), \ v_t = 1 - \alpha_t$$

Denoising step: $t \in [0, 1]$, $\mu_0 = p_T \approx N(0, I)$, $\rightarrow \mu_1 \approx p_0$,

$$\mathrm{d}\mathbf{v}_t = \frac{1}{2}\beta(t)\mathbf{v}_t\mathrm{d}t + \beta(t)\nabla_{\mathbf{v}}\log p_{T-t}(\mathbf{v}_t)\mathrm{d}t + \sqrt{\beta(t)}\mathrm{d}\mathbf{w}_t, \ \mathbf{v}_0 \sim \mu_0.$$



Conditional diffusion modelling: Sample from

 $p_0(x_0|y) \propto p(y|x_0)p_0(x_0),$ $-\log p(y|x) = \frac{1}{2}(y - Hv)^{\mathrm{T}}R^{-1}(Hv - y) + \text{const.}$



Conditional diffusion modelling: Sample from

 $p_0(x_0|y) \propto p(y|x_0)p_0(x_0),$ $-\log p(y|x) = \frac{1}{2}(y - Hv)^{\mathrm{T}}R^{-1}(Hv - y) + \text{const.}$

Prediction step. $\mu_0 = N(0, I) \rightarrow v_1 \sim \hat{\mu}_1$:

$$\mathrm{d}\mathbf{v}_t = \frac{1}{2}\beta(t)\mathbf{v}_t\mathrm{d}t + \beta(t)\nabla_{\mathbf{v}}\log p_{T-t}(\mathbf{v}_t)\mathrm{d}t + \sqrt{\beta(t)}\mathrm{d}\mathbf{w}_t$$

Analysis step. $\hat{\mu}_1 = p_0 \rightarrow \mu_1 = p_0(\cdot|y)$.



• All examples so far introduced two sequences of PDFs $\{\mu_n\}$ and $\{\hat{\mu}_n\}$, for $n \ge 0$. (Conditional diffusion model n = 0, 1.)



- All examples so far introduced two sequences of PDFs $\{\mu_n\}$ and $\{\hat{\mu}_n\}$, for $n \ge 0$. (Conditional diffusion model n = 0, 1.)
- Those marginal PDFs do not define the desired stochastic processes in $\{v_n\}$ and $\{\hat{v}_n\}$ uniquely.



- All examples so far introduced two sequences of PDFs $\{\mu_n\}$ and $\{\hat{\mu}_n\}$, for $n \ge 0$. (Conditional diffusion model n = 0, 1.)
- Those marginal PDFs do not define the desired stochastic processes in $\{v_n\}$ and $\{\hat{v}_n\}$ uniquely.
- This conversion requires joint distributions.



- All examples so far introduced two sequences of PDFs $\{\mu_n\}$ and $\{\hat{\mu}_n\}$, for $n \ge 0$. (Conditional diffusion model n = 0, 1.)
- Those marginal PDFs do not define the desired stochastic processes in $\{v_n\}$ and $\{\hat{v}_n\}$ uniquely.
- This conversion requires joint distributions.
- Consider mean-field processes of type

 $\begin{array}{ll} \textbf{Prediction:} & \hat{v}_{n+1} = \Phi_n(v_n, \mu_n, \xi_n),\\ \textbf{Analysis:} & v_{n+1} = \hat{\Phi}_n(\hat{v}_{n+1}, \hat{\mu}_{n+1}, \hat{\xi}_n), \end{array}$

for appropriate functions Φ_n , $\hat{\Phi}_n$; i.i.d. random variables ξ_n , $\hat{\xi}_n$; μ_n the law of v_n and $\hat{\mu}_n$ the law of \hat{v}_n .



Monte Carlo/particle implementations:

• Introduce J particles $v_{n,j}$ and $\hat{v}_{n,j}$, j = 1, ..., J, and their associated (random) empirical measures

$$\mu_n^J(v) = \frac{1}{J} \sum_{j=1}^M \delta(v - v_{n,j}), \qquad \hat{\mu}_n^J(v) = \frac{1}{J} \sum_{j=1}^M \delta(v - \hat{v}_{n,j})$$

at each iteration index $n \ge 0$.



Monte Carlo/particle implementations:

• Introduce J particles $v_{n,j}$ and $\hat{v}_{n,j}$, j = 1, ..., J, and their associated (random) empirical measures

$$\mu_n^J(v) = \frac{1}{J} \sum_{j=1}^M \delta(v - v_{n,j}), \qquad \hat{\mu}_n^J(v) = \frac{1}{J} \sum_{j=1}^M \delta(v - \hat{v}_{n,j})$$

at each iteration index $n \ge 0$.

• Interacting particles update step: $n \geq 0$, $j = 1, \ldots, M$

with (generally) approximate maps Φ_n^a and $\hat{\Phi}_n^a$.



Theory:

- Long time behaviour of μ_n and $\hat{\mu}_n$ as $n \to \infty$.
- Dependence of μ_n and $\hat{\mu}_n$ on initial μ_0 as $n \to \infty$.
- Dependence of μ_n and $\hat{\mu}_n$ on data $\{y_k^{\dagger}\}_{k=1}^n$.

• ...



Theory:

- Long time behaviour of μ_n and $\hat{\mu}_n$ as $n \to \infty$.
- Dependence of μ_n and $\hat{\mu}_n$ on initial μ_0 as $n \to \infty$.
- Dependence of μ_n and $\hat{\mu}_n$ on data $\{y_k^{\dagger}\}_{k=1}^n$.

• ...

Algorithms:

- Two types of approximation errors:
 - Monte Carlo sampling error
 - Approximate prediction and analysis steps $\Phi_n^{\rm a}$ and $\hat{\Phi}_n^{\rm a}$, respectively.
- Error propagation and stability as $n \to \infty$.

• ...



Theory:

- Long time behaviour of μ_n and $\hat{\mu}_n$ as $n \to \infty$.
- Dependence of μ_n and $\hat{\mu}_n$ on initial μ_0 as $n \to \infty$.
- Dependence of μ_n and $\hat{\mu}_n$ on data $\{y_k^{\dagger}\}_{k=1}^n$.

• ...

Algorithms:

- Two types of approximation errors:
 - Monte Carlo sampling error
 - Approximate prediction and analysis steps $\Phi_n^{\rm a}$ and $\hat{\Phi}_n^{\rm a}$, respectively.
- Error propagation and stability as $n \to \infty$.

• ...

Remark. Related but also fundamentally different to sequential Monte-Carlo methods.

UQ & DA, UP & SFB 1294

Content



- Part I. Introduction to data assimilation (DA)
 - What is DA?
 - Examples
 - Applications and challenges
- Part II. The ensemble Kalman filter (EnKF)
 - Standard Kalman filter
 - Mean-field EnKF formulation & Monte Carlo implementation
 - Why the EnKF?
 - Why not?
- Part III. Control & DA
 - Coupling of measures
 - Homotopy and data-driven control

<u>Reference</u>: Ensemble Kalman Methods: A Mean Field Perspective; Edoardo Calvello, SR, Andrew M. Stuart, arXiv:2209.11371.

UQ & DA, UP & SFB 1294



Stochastic dynamical system:

$$\mathbf{v}_{n+1} = F\mathbf{v}_n + \mathbf{b} + B\xi_n,$$

 $\xi_n \sim N(0, I)$ i.i.d.



Stochastic dynamical system:

$$\mathbf{v}_{n+1} = F\mathbf{v}_n + \mathbf{b} + B\xi_n,$$

 $\xi_n \sim N(0, I)$ i.i.d.

Observation model:

$$y = H\hat{v} + R^{1/2}\eta,$$

 $\eta \sim N(0, I)$ and prior distribution $\hat{v} \sim N(\hat{m}^{v}, \hat{C}^{vv})$.



Gaussian DA:

$$v_n \sim \mathrm{N}(m_n^v, C_n^{vv}), \qquad \hat{v}_n \sim \mathrm{N}(\hat{m}_n^v, \hat{C}_n^{vv}).$$



Gaussian DA:

$$v_n \sim \mathrm{N}(m_n^v, C_n^{vv}), \qquad \hat{v}_n \sim \mathrm{N}(\hat{m}_n^v, \hat{C}_n^{vv}).$$

Prediction step:

$$\begin{split} \hat{m}_{n+1}^{v} &= Fm_{n}^{v} + b, \\ \hat{C}_{n+1}^{vv} &= FC_{n}^{vv}F^{\mathrm{T}} + BB^{\mathrm{T}}. \end{split}$$



Gaussian DA:

$$v_n \sim \mathrm{N}(m_n^v, C_n^{vv}), \qquad \hat{v}_n \sim \mathrm{N}(\hat{m}_n^v, \hat{C}_n^{vv}).$$

Prediction step:

$$\begin{split} \hat{m}_{n+1}^{v} &= Fm_{n}^{v} + b, \\ \hat{C}_{n+1}^{vv} &= FC_{n}^{vv}F^{\mathrm{T}} + BB^{\mathrm{T}}. \end{split}$$

Analysis step: Given data y_{n+1}^{\dagger}

$$m_{n+1}^{v} = \hat{m}_{n+1}^{v} + K_{n+1}(y_{n+1}^{\dagger} - H\hat{m}_{n+1}^{v}),$$

$$C_{n+1}^{vv} = \hat{C}_{n+1}^{vv} - K_{n+1}H\hat{C}_{n+1}^{vv}.$$

with Kalman gain

$$K_{n+1} = \hat{C}_{n+1}^{vy} (\hat{C}_n^{yy})^{-1} = \hat{C}_{n+1}^{vv} H^{\mathrm{T}} (H\hat{C}_{n+1}^{vv} H^{\mathrm{T}} + R)^{-1}.$$


Gaussian mean-field DA:

$$v_n o \hat{v}_{n+1}, \qquad \hat{v}_{n+1} o v_{n+1}.$$



Gaussian mean-field DA:

$$v_n \rightarrow \hat{v}_{n+1}, \qquad \hat{v}_{n+1} \rightarrow v_{n+1}.$$

Prediction & output step: $v_n \sim \mu_n \& \xi_n$, η_n i.i.d. N(0, *I*)

 $\hat{v}_{n+1} = Fv_n + b + B\xi_n,$ $\hat{y}_{n+1} = H\hat{v}_{n+1} + R^{1/2}\eta_n.$



Gaussian mean-field DA:

$$v_n \to \hat{v}_{n+1}, \qquad \hat{v}_{n+1} \to v_{n+1}.$$

Prediction & output step: $v_n \sim \mu_n \& \xi_n, \eta_n \text{ i.i.d. N}(0, I)$

$$\hat{v}_{n+1} = Fv_n + b + B\xi_n,$$

 $\hat{y}_{n+1} = H\hat{v}_{n+1} + R^{1/2}\eta_n.$

Analysis step: Given data y_{n+1}^{\dagger} , linear mean-field update

$$v_{n+1} = \hat{v}_{n+1} + K_{n+1}(y_{n+1}^{\dagger} - \hat{y}_{n+1})$$

with Kalman gain

$$K_{n+1} = \hat{C}_{n+1}^{vy} \left(\hat{C}_{n+1}^{yy} \right)^{-1} = \hat{C}_{n+1} H^{\mathrm{T}} (H \hat{C}_{n+1} H^{\mathrm{T}} + R)^{-1}.$$



EnKF:
$$J$$
 particles $v_{n,j}$, $\hat{v}_{n,j}$, $j = 1, ..., J$, empirical estimators
 $\mathbf{m}_n^{\mathbf{v}} := \mu_n^J[\hat{\mathbf{v}}] = \mathbb{E}_n^J[\hat{\mathbf{v}}], \ \hat{\mathbf{C}}_n^{\mathbf{vy}} := \mathbb{E}_n^J \left[(\hat{\mathbf{v}} - \mathbb{E}_n^J[\hat{\mathbf{v}}])(\hat{\mathbf{y}} - \mathbb{E}_n^J[\hat{\mathbf{y}}])^{\mathrm{T}} \right], \text{ etc.}$



EnKF:
$$J$$
 particles $v_{n,j}$, $\hat{v}_{n,j}$, $j = 1, ..., J$, empirical estimators
 $\mathbf{m}_n^{\mathbf{v}} := \mu_n^J[\hat{v}] = \mathbb{E}_n^J[\hat{v}], \ \hat{\mathbf{C}}_n^{\mathbf{v}\mathbf{y}} := \mathbb{E}_n^J\left[(\hat{v} - \mathbb{E}_n^J[\hat{v}])(\hat{y} - \mathbb{E}_n^J[\hat{y}])^{\mathrm{T}}\right], \text{ etc.}$

Prediction & output step: j = 1, ..., J, $\xi_{n,j}$, $\eta_{n,j}$ i.i.d. N(0, *I*)

$$\hat{v}_{n+1,j} = F v_{n,j} + b + B^{\mathrm{T}} \xi_{n,j},$$

 $\hat{y}_{n+1,j} = H \hat{v}_{n+1,j} + R^{1/2} \eta_{n,j}.$



EnKF:
$$J$$
 particles $v_{n,j}$, $\hat{v}_{n,j}$, $j = 1, ..., J$, empirical estimators
 $\mathbf{m}_n^{\mathbf{v}} := \mu_n^J[\hat{v}] = \mathbb{E}_n^J[\hat{v}], \ \hat{\mathbf{C}}_n^{\mathbf{v}\mathbf{y}} := \mathbb{E}_n^J\left[(\hat{v} - \mathbb{E}_n^J[\hat{v}])(\hat{y} - \mathbb{E}_n^J[\hat{y}])^{\mathrm{T}}\right], \text{ etc.}$

Prediction & output step: j = 1, ..., J, $\xi_{n,j}$, $\eta_{n,j}$ i.i.d. N(0, *I*) $\hat{v}_{n+1,j} = F v_{n,j} + b + B^{T} \xi_{n,j}$, $\hat{y}_{n+1,j} = H \hat{v}_{n+1,j} + R^{1/2} \eta_{n,j}$.

EnKF analysis step: Given data y_{n+1}^{\dagger} , j = 1, ..., J, $v_{n+1,j} = \hat{v}_{n+1,j} + K_{n+1}(y_{n+1}^{\dagger} - \hat{y}_{n+1,j})$ with (empirical) Kalman gain

$$\mathbf{K}_{n+1} := \hat{\mathbf{C}}_{n+1}^{vy} \left(\hat{\mathbf{C}}_{n+1}^{yy} \right)^{-1}.$$



• Often used with $J \ll d$ when $d \gg 1$ (model reduction). Additional computational hacks required (localisation, inflation).



- Often used with J ≪ d when d ≫ 1 (model reduction). Additional computational hacks required (localisation, inflation).
- Can be extended to nonlinear DA problems.
 - Prediction & output step:

$$\hat{v}_{n+1} = \Phi(v_n, \xi_n),$$

 $\hat{y}_{n+1} = h(\hat{v}_{n+1}) + R^{1/2}\eta_n.$

• EnKF analysis step: unaltered.



- Often used with $J \ll d$ when $d \gg 1$ (model reduction). Additional computational hacks required (localisation, inflation).
- Can be extended to nonlinear DA problems.
 - Prediction & output step:

$$\hat{v}_{n+1} = \Phi(v_n, \xi_n),$$

 $\hat{y}_{n+1} = h(\hat{v}_{n+1}) + R^{1/2}\eta_n.$

- EnKF analysis step: unaltered.
- The EnKF is <u>affine invariant</u> (like Newton's method) and <u>derivative-free</u>.



- Often used with $J \ll d$ when $d \gg 1$ (model reduction). Additional computational hacks required (localisation, inflation).
- Can be extended to nonlinear DA problems.
 - Prediction & output step:

$$\hat{v}_{n+1} = \Phi(v_n, \xi_n),$$

 $\hat{y}_{n+1} = h(\hat{v}_{n+1}) + R^{1/2}\eta_n.$

- EnKF analysis step: unaltered.
- The EnKF is <u>affine invariant</u> (like Newton's method) and <u>derivative-free</u>.
- But why like this? Best Unbiased Linear Estimator (BLUE) .



Prediction step: $(\hat{v}_{n+1}, \hat{y}_{n+1})$ with potentially unknown distribution (generative model)

Best linear (point) estimator for \hat{v}_{n+1} given \hat{y}_{n+1} :

$$e(y) = b + Ky$$

with

$$(b, K) = \arg \min \mathbb{E} \left[\| \hat{v}_{n+1} - e(\hat{y}_{n+1}) \|^2 \right].$$

Plug-in estimator: Given samples $\{(\hat{v}_{n,j}, \hat{y}_{n,j})\}$

$$\mathbf{K} = \hat{\mathbf{C}}_{n+1}^{vy} (\hat{\mathbf{C}}_{n+1}^{yy})^{-1}$$
$$\mathbf{b} = \hat{\mathbf{m}}_{n+1}^{v} - \mathbf{K}\hat{\mathbf{m}}_{n+1}^{y}$$

Remark. Corresponds to update of the ensemble mean in an EnKF.



Data generation: v_{n+1}^{\dagger} true value (fixed),

$$y_{n+1}^{\dagger} = Hv_{n+1}^{\dagger} + R^{1/2}\eta_{n+1}^{\dagger}, \qquad \eta_{n+1}^{\dagger} \sim N(0, I).$$



Data generation: v_{n+1}^{\dagger} true value (fixed), $y_{n+1}^{\dagger} = Hv_{n+1}^{\dagger} + R^{1/2}\eta_{n+1}^{\dagger}, \qquad \eta_{n+1}^{\dagger} \sim N(0, I).$

Bias of BLUE:

$$\mathbb{E}^{\dagger}\left[e(y_{n+1}^{\dagger})-v_{n+1}^{\dagger}\right]=(I-\mathcal{K}H)(\hat{m}_{n+1}^{v}-v_{n+1}^{\dagger}).$$

Covariance of BLUE:

 $\operatorname{var}^{\dagger}[e(y_{n+1}^{\dagger})] = KRK^{\mathrm{T}}.$



Data generation: v_{n+1}^{\dagger} true value (fixed),

$$y_{n+1}^{\dagger} = Hv_{n+1}^{\dagger} + R^{1/2}\eta_{n+1}^{\dagger}, \qquad \eta_{n+1}^{\dagger} \sim N(0, I)$$

Bias of BLUE:

$$\mathbb{E}^{\dagger}\left[e(y_{n+1}^{\dagger})-v_{n+1}^{\dagger}\right]=(I-\mathcal{K}H)(\hat{m}_{n+1}^{v}-v_{n+1}^{\dagger}).$$

Covariance of BLUE:

 $\operatorname{var}^{\dagger}[e(y_{n+1}^{\dagger})] = KRK^{\mathrm{T}}.$

Remark. Classical bias-variance trade-off of frequentist uncertainty quantification (UQ).



Challenge: In DA we need more than an estimator. We also need to propagate its uncertainty!



Challenge: In DA we need more than an estimator. We also need to propagate its uncertainty!

Solution: Introduce random variable v_{n+1} such that

$$\begin{split} m_{n+1}^{\mathsf{v}} &= e(y_{n+1}^{\dagger}), \\ C_{n+1}^{\mathsf{vv}} &\geq \mathsf{KR}\mathsf{K}^{\mathrm{T}}, \\ C_{n+1}^{\mathsf{vv}} &\leq \hat{C}_{n+1}^{\mathsf{vv}}. \end{split}$$

Mean-field EnKF:

$$\hat{y}_{n+1} = H\hat{v}_{n+1} + R^{1/2}\eta_{n+1},$$

$$v_{n+1} = \hat{v}_{n+1} + K(y_{n+1}^{\dagger} - \hat{y}_{n+1}).$$

since

$$C_{n+1}^{vv} = (I - KH)\hat{C}_{n+1}^{vv}(I - KH)^{\mathrm{T}} + KRK^{\mathrm{T}} = (I - KH)\hat{C}_{n+1}^{vv}.$$



Drift-diffusion model

$$\mathrm{d}\mathbf{v}_t = f(\mathbf{v}_t)\mathrm{d}t + (2\sigma)^{1/2}\mathrm{d}\mathbf{w}_t$$

 $\sigma > 0$, w_t Brownian motion.



Drift-diffusion model

$$\mathrm{d}\mathbf{v}_t = f(\mathbf{v}_t)\mathrm{d}t + (2\sigma)^{1/2}\mathrm{d}\mathbf{w}_t$$

 $\sigma > 0$, w_t Brownian motion.

Given:

i) Initial conditions

 $v_0 \sim \mu_0$,

ii) Observation at time t = T > 0

 $y_T = Hv_T + \eta_T, \qquad \eta_T \sim \mathrm{N}(0, R),$

H the (linear) forward model.



<u>Two-dimensional nonlinear diffusion</u>: $v = (v_1, v_2)^{\top}$. Drift term:

$$f(v) = -\nabla V(v), \quad V(v) = \frac{\lambda_1}{2} \left(v_2 - 2 + \beta(v_1)^2\right)^2 + \frac{\lambda_2}{2} \left(\frac{(v_1)^4}{2} - (v_1)^2\right)$$

with parameters $\lambda_1=$ 2000, $\lambda_2=$ 5, and $\beta=1/5,$ diffusion constant $\sigma=1.$



<u>Two-dimensional nonlinear diffusion</u>: $v = (v_1, v_2)^{\top}$. Drift term:

$$f(\mathbf{v}) = -\nabla V(\mathbf{v}), \quad V(\mathbf{v}) = \frac{\lambda_1}{2} \left(\mathbf{v}_2 - 2 + \beta (\mathbf{v}_1)^2 \right)^2 + \frac{\lambda_2}{2} \left(\frac{(\mathbf{v}_1)^4}{2} - (\mathbf{v}_1)^2 \right)$$

with parameters $\lambda_1=$ 2000, $\lambda_2=$ 5, and $\beta=1/5,$ diffusion constant $\sigma=1.$

Initial distribution:

$$v_{1,0} \sim N(1.5, 0.0625), \quad v_{2,0} = 2 - \beta v_{1,0}^2.$$

Observation:

$$H = (1 \ 0), \quad T = 1, \quad R = 0.01, \quad y_T^{\dagger} = -1.5.$$







Drift-diffusion model

 $\mathrm{d}\mathbf{v}_t = f(\mathbf{v}_t)\mathrm{d}t + g_t(\mathbf{v}_t)\mathrm{d}t + (2\sigma)^{1/2}\mathrm{d}\mathbf{w}_t.$



Figure: Left panel: Initial and final particle positions under the controlled evolution process. Right panel: Particle positions at intermediate times

Content



- Part I. Introduction to data assimilation (DA)
 - What is DA?
 - Examples
 - Applications and challenges
- Part II. The ensemble Kalman filter (EnKF)
 - Standard Kalman filter
 - Mean-field EnKF formulation & Monte Carlo implementation
 - Why the EnKF & why not?
- Part III. Control & DA
 - Coupling of measures
 - Homotopy and data-driven control

References:

SR, Data assimilation: A dynamic homotopy-based coupling approach, arXiv:2209.05279, to appear STUOD Proceedings, 2023. Chen, Y. et al, Gradient Flows for Sampling: Mean-Field Models, Gaussian Approximations and Affine Invariance, arXiv:2302.11024.



<u>Prior</u> (prediction): Random variable \hat{v} with

 $\hat{\mathbf{v}} \sim \hat{\mu}$

negative log-likelihood:

nonl. regression:
$$l(v) = \frac{1}{2}(h(v) - y^{\dagger})^{\mathrm{T}}R^{-1}(h(v) - y^{\dagger}),$$

h nonlinear forward map, R error covariance matrix, $y^{\dagger} \in \mathbb{R}^{N_y}$ the data.

Bayesian posterior (analysis):

$$\mu(\mathbf{v}) := \frac{e^{-l(\mathbf{v})}\,\hat{\mu}(\mathbf{v})}{\hat{\mu}\left[e^{-l}\right]}.$$



Find joint distribution π with desired marginals, i.e.,

 $(\hat{\mathbf{v}},\mathbf{v})\sim\pi$

such that

$$\pi(\hat{\mathbf{v}},\mathbf{v}) = \pi(\mathbf{v}|\hat{\mathbf{v}})\,\hat{\mu}(\hat{\mathbf{v}}) = \pi(\hat{\mathbf{v}}|\mathbf{v})\mu(\mathbf{v}).$$

The set of all such joint distributions is denoted by Π .

(1)



Find joint distribution π with desired marginals, i.e.,

 $(\hat{\mathbf{v}},\mathbf{v})\sim\pi$

such that

$$\pi(\hat{v}, \mathbf{v}) = \pi(\mathbf{v}|\hat{v})\,\hat{\mu}(\hat{v}) = \pi(\hat{v}|\mathbf{v})\mu(\mathbf{v}).$$

The set of all such joint distributions is denoted by Π .

Catch: Π is not empty: $\pi(\hat{v}, v) = \hat{\mu}(\hat{v})\mu(v) \in \Pi$.

Selection of a unique π^* via:

- optimal transportation (minimise expected distance between \hat{v} and v)
- Schrödinger bridges (minimise the Kullback–Leibler divergence to some reference measure)
- homotopy (Fisher-Rao gradient flow)

(1)



Idea:¹ Define transport map T as the time-one flow map of a (mean-field) ODE

$$rac{\mathrm{d}}{\mathrm{d} au} v_{ au} = f_{ au}(v_{ au}), \qquad v_0 \sim \hat{\mu}, \quad au \in [0, 1].$$

Liouville equation for $v_{\tau} \sim \pi_{\tau}$:

$$\partial_{\tau}\pi_{\tau}=-\nabla\cdot\left(\pi_{\tau}f_{\tau}\right)$$

Problem: How to fix f_{τ} appropriately given $\pi_0 = \hat{\mu}$ and $\pi_1 = \mu$?

¹Daum & Huang, ICASSP, 2011; SR, BIT, 2011, see also normalising flows in ML



Homotopy Bayes:

$$\pi_{\tau}(\mathbf{v}) \propto e^{-\tau l(\mathbf{v})} \hat{\mu}(\mathbf{v}).$$

implying

$$\partial_{\tau}\pi_{\tau} = -\pi_{\tau}\left(I - \pi_{\tau}[I]\right)$$

Fisher-Rao gradient flow:

$$\partial_ au \pi_ au = -\mathsf{grad}\,\mathcal{E}(\pi_ au)$$

with functional

$$\mathcal{E}(\pi) = \int I(\mathbf{v})\pi(\mathbf{v})\mathrm{d}\mathbf{v}.$$



Mean-field equations in the parameters $v \in \mathbb{R}^d$:

$$\begin{aligned} \frac{\mathrm{d}}{\mathrm{d}\tau} \mathbf{v}_{\tau} &= f(\mathbf{v}_{\tau}; \pi_{\tau}) \\ \nabla \cdot (\pi_{\tau} f) &= \operatorname{grad} \mathcal{E}(\pi_{\tau}), \end{aligned}$$

where π_{τ} denotes the law of v_{τ} and $v_0 \sim \hat{\mu}$.

Remark: Mean-field drift term $f(v; \pi)$ is **not** uniquely determined.



Mean-field equations in the parameters $v \in \mathbb{R}^d$:

$$\frac{\mathrm{d}}{\mathrm{d}\tau} \mathbf{v}_{\tau} = f(\mathbf{v}_{\tau}; \pi_{\tau})$$
$$\nabla \cdot (\pi_{\tau} f) = \operatorname{grad} \mathcal{E}(\pi_{\tau}),$$

where π_{τ} denotes the law of v_{τ} and $v_0 \sim \hat{\mu}$.

Remark: Mean-field drift term $f(v; \pi)$ is **not** uniquely determined.

<u>Affine invariant</u> choices: Find potential function ψ such that

$$f(v;\pi) = (\nabla^2 \psi)^{-1} \nabla \psi(v;\pi)$$
 or $f(v;\pi) = C^{vv} \nabla \psi(v;\pi)$

with <u>covariance matrix</u> $C^{\nu\nu}$.



Drift-diffusion model

$$\mathrm{d}\mathbf{v}_t = f(\mathbf{v}_t)\mathrm{d}t + (2\sigma)^{1/2}\mathrm{d}\mathbf{w}_t$$

 $\sigma > 0$, w_t Brownian motion.



Drift-diffusion model

$$\mathrm{d}\mathbf{v}_t = f(\mathbf{v}_t)\mathrm{d}t + (2\sigma)^{1/2}\mathrm{d}\mathbf{w}_t$$

 $\sigma > 0$, w_t Brownian motion.

Given:

i) Initial conditions

 $v_0 \sim \mu_0$,

ii) Observations at time t = T > 0

 $y_T = Hv_T + \eta_T, \qquad \eta_T \sim N(0, R),$

H the (linear) forward model.



Controlled SDE

$$\mathrm{d}\tilde{v}_t = f(\tilde{v}_t)\mathrm{d}t + \frac{g_t(\tilde{v}_t)\mathrm{d}t}{g_t(\tilde{v}_t)\mathrm{d}t} + \sqrt{2\sigma}\mathrm{d}w_t$$

for appropriate control $g_t(v)$, $\tilde{v}_t \sim \tilde{\mu}_t$.



Controlled SDE

$$\mathrm{d}\tilde{v}_t = f(\tilde{v}_t)\mathrm{d}t + g_t(\tilde{v}_t)\mathrm{d}t + \sqrt{2\sigma}\mathrm{d}w_t$$

for appropriate control $g_t(v)$, $\tilde{v}_t \sim \tilde{\mu}_t$.

Schrödinger bridge problem: Find control g_t minimizing Kullback-Leibler divergence

$$\mathrm{KL}(\hat{\mu}|\tilde{\mu}) = \int_0^T \int_{\mathbb{R}^d} \hat{\mu}_t(v) (\log \hat{\mu}_t(v) - \log \tilde{\mu}_t(v)) \mathrm{d}v \, \mathrm{d}t$$

subject to

$$\widetilde{\mu}_0 = \mu_0, \qquad \widetilde{\mu}_T = \mu_T.$$



Homotopy:

$$\tilde{\mu}_t(\mathbf{v}) = Z_t^{-1} e^{-\frac{t}{T}I(\mathbf{v})} \hat{\mu}_t(\mathbf{v})$$

with $Z_t = \int e^{-\frac{t}{T}l(v)}\hat{\mu}_t(v) dx$. Obviously

 $\tilde{\mu}_0 = \mu_0, \qquad \tilde{\mu}_T = \mu_T.$



Homotopy:

$$\tilde{\mu}_t(\mathbf{v}) = Z_t^{-1} e^{-\frac{t}{T}I(\mathbf{v})} \hat{\mu}_t(\mathbf{v})$$

with $Z_t = \int e^{-\frac{t}{T}I(v)}\hat{\mu}_t(v) dx$. Obviously

 $\tilde{\mu}_0 = \mu_0, \qquad \tilde{\mu}_T = \mu_T.$

Given $v_t \sim \hat{\mu}_t$:

$$\mathrm{d}\mathbf{v}_t = f(\mathbf{v}_t)\mathrm{d}t + (2\sigma)^{1/2}\mathrm{d}\mathbf{w}_t, \qquad \mathbf{v}_0 \sim \mu_0$$

Desired $v_t^{\rm h} \sim \pi_t^{\rm h}$:

$$\mathrm{d}\tilde{v}_t = f(\tilde{v}_t)\mathrm{d}t - \frac{\sigma t}{T} \nabla I(\tilde{v}_t)\mathrm{d}t + g_t(\tilde{v}_t)\mathrm{d}t + \sqrt{2\sigma}\mathrm{d}w_t$$

for appropriate control $g_t(v)$.


Desired $\tilde{v}_t \sim \tilde{\mu}_t$:

$$\mathrm{d}\tilde{v}_t = f(\tilde{v}_t)\mathrm{d}t - \frac{2\sigma t}{T}\nabla I(\tilde{v}_t)\mathrm{d}t + \sqrt{2\sigma}\mathrm{d}w_t + g_t(\tilde{v}_t)\mathrm{d}t$$

for appropriate control $g_t(v)$.

Control g_t satisfies PDE

$$\nabla \cdot (\tilde{\mu}_t \mathbf{g}_t) = \frac{1}{T} \tilde{\mu}_t \left(\mathbf{I} + t \nabla \mathbf{I} \cdot \mathbf{f} - \frac{\sigma t^2}{T} \|\nabla \mathbf{I}\|^2 \right) + \tilde{\mu}_t \frac{\dot{Z}_t}{Z_t}.$$



Constant gain approximation (EnKF-like):

$$g_t^{\text{KF}}(v) = -K_t \left(\frac{1}{2} \left(Hv + Hm_t \right) - y_T^{\dagger} \right) \\ - \frac{t}{2T} C_t^{vv} H^{\top} R^{-1} H \left(f(x) + \tilde{\mu}_t[f] \right)$$

with Kalman gain

$$\mathcal{K}_t = \left\{ \frac{1}{T} C_t^{\mathsf{vv}} + \frac{t}{T} C_t^{\mathsf{vf}} - \frac{2\sigma t^2}{T^2} C_t^{\mathsf{vv}} H^\top R^{-1} H \right\} H^\top R^{-1}.$$





Nonlinear example II





Figure: Left panel: Initial and final particle positions under the controlled evolution process. Right panel: Particle positions at intermediate times $t_k \in [0, 1]$.



$$\frac{\mathrm{d}}{\mathrm{d}t}X_t = f(X_t), \qquad f(x, y, z) = \begin{pmatrix} a(y-x) \\ x(b-z) - y \\ xy - cz \end{pmatrix}$$
(2)

with parameters a = 10, b = 28 and c = 8/3. The first component is observed.

$M/\Delta t_{ m obs}$	5	10	15
0.05	0.5712/0.5457	0.5620/0.5475	0.5659/0.5496
0.10	0.8466/0.7735	0.8171/0.7627	0.8229/0.7707
0.12	0.9606/0.8645	0.9515/0.8621	0.9375/0.8615

Table: RMSE for the **ensemble square root filter** and the **homotopy approach** in terms of ensemble sizes $M \in \{5, 10, 15\}$ and observation intervals $\Delta t_{\rm obs} \in \{0.05, 0.1, 0.12\}$.



Prediction step. $\mu_0 = N(0, I) \rightarrow v_1 \sim \hat{\mu}_1$:

$$dv_t = \frac{1}{2}\beta(t)v_t dt + \beta(t)\nabla_v \log p_{T-t}(v_t)dt + \sqrt{\beta(t)}dw_t$$
$$= f_t(v_t)dt + \sqrt{\beta(t)}dw_t$$

Conditioning.

$$I(v_1) = \frac{1}{2}(Hv_1 - y)^{\mathrm{T}}R^{-1}(Hv_1 - y).$$



Prediction step. $\mu_0 = N(0, I) \rightarrow v_1 \sim \hat{\mu}_1$:

$$dv_t = \frac{1}{2}\beta(t)v_t dt + \beta(t)\nabla_v \log p_{T-t}(v_t)dt + \sqrt{\beta(t)}dw_t$$
$$= f_t(v_t)dt + \sqrt{\beta(t)}dw_t$$

Conditioning.

$$I(v_1) = \frac{1}{2}(Hv_1 - y)^{\mathrm{T}}R^{-1}(Hv_1 - y).$$

Controlled diffusion.

 $\mathrm{d}\tilde{v}_t = f_t(\tilde{v}_t)\mathrm{d}t - \beta(t)tH^{\mathrm{T}}R^{-1}(H\tilde{v}_t - y) + g_t(\tilde{v}_t)\mathrm{d}t + \sqrt{\beta(t)}\mathrm{d}w_t$

with $2\sigma \rightarrow \beta(t)$ and T = 1.



Pros:

- \bullet no initial approximation of $\mu_{\mathcal{T}}$ needed
- no forward-backward iterations required
- avoids high-dimensional optimisation problem
- derivative-free
- applicable to conditional diffusion modelling



Pros:

- \bullet no initial approximation of $\mu_{\mathcal{T}}$ needed
- no forward-backward iterations required
- avoids high-dimensional optimisation problem
- derivative-free
- applicable to conditional diffusion modelling

Cons:

- $\bullet\,$ numerical approximation of $\tilde{\mu}_t$ and of an elliptic PDE required
- EnKF-type implementation requires that $\tilde{\mu}_t$ remains close to Gaussian
- dynamics of the interacting particle system not understood



THE END

UQ & DA, UP & SFB 1294