# Bayesian Inversion: Algorithms

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A.M. Stuart. *Inverse problems: a Bayesian perspective.* Acta Numerica **19**(2010).

~masdr/BOOKCHAPTERS/stuart15c.pdf

- M. Dashti, K.J.H. Law, A.M. Stuart and J. Voss. MAP estimators and posterior consistency . . . .
  Inverse Problems, 29(2013), 095017. arxiv:1303.4795.
- F. Pinski, G. Simpson, A.M. Stuart and H. Weber. Kullback-Leibler approximation for probability measures on infinite dimensional spaces. In preparation.
- S.L. Cotter, G.O. Roberts, A.M. Stuart and D. White. *MCMC methods for functions*.... Statistical Science **28**(2013). arxiv:1202.0709.
- M. Hairer, A.M.Stuart and S. Vollmer. Spectral gaps for a Metropolis-Hastings algorithm . . . . arxiv: 1112.1392.

#### Outline

- SETTING AND ASSUMPTIONS
- MAP ESTIMATORS
- 3 KULLBACK-LEIBLER APPROXIMATION
- SAMPLING
- **5** CONCLUSIONS



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# The Setting

- Probability measure  $\mu$  on Hilbert space H.
- Reference measure  $\mu_0$  (often a prior).
- $\mu$  related to  $\mu_0$  by (often Bayes' Theorem)

$$\frac{d\mu}{d\mu_0}(u) = \frac{1}{Z_\mu} \exp\left(-\Phi(u)\right).$$

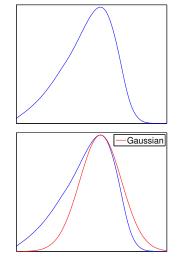
• Another way of saying the same thing:

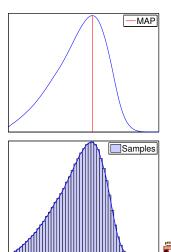
$$\mathbb{E}^{\mu}f(u) = \frac{1}{Z_{\mu}}\mathbb{E}^{\mu_0}\Big(\exp(-\Phi(u))f(u)\Big).$$

• How do we get information from  $\mu$  if we know  $\mu_0$  and  $\Phi$ ?



#### The Talk In One Picture







# The Assumptions

- $\mu_0 = N(0, C_0)$  a centred Gaussian measure on H.
- $\mu_0(X) = 1$ ; X (Banach) continuously embedded in H.
- Let  $E = \mathcal{D}(C_0^{-\frac{1}{2}})$  (Cameron-Martin space).
- Then  $E \subset X \subseteq H$ . E (Hilbert) compactly embedded in X.
- The function  $\Phi \in C(X; \mathbb{R}^+)$ .
- For all u, v with  $||u||_X \le r$ ,  $||v||_X \le r$  there are  $M_i(r)$ :

$$|\Phi(u)| \le M_1(r);$$
  
 $|\Phi(u) - \Phi(v)| \le M_2(r)||u - v||.$ 



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# Probability Maximizers and Tikhonov Regularization

Define the Tikhonov-regularized LSQ functional  $I: E \to \mathbb{R}^+$  by

$$I(u) := \frac{1}{2} \|C_0^{-\frac{1}{2}}u\|^2 + \Phi(u).$$

Let  $B^{\delta}(z)$  be a ball of radius  $\delta$  in X centred at  $z \in E = \mathcal{D}(C_0^{-\frac{1}{2}})$ .

#### Theorem

(Dashti, Law, S and Voss, 2013). The probability measure  $\mu$  and functional I are related by

$$\lim_{\delta \to 0} \frac{\mu(B^{\delta}(z_1))}{\mu(B^{\delta}(z_2))} = \exp(I(z_2) - I(z_1)).$$

Thus probability maximizers are minimizers of the regularized Tikhonov functional *I*.

# Existence of Probability Maximizers

The minimization is well-defined:

#### Theorem

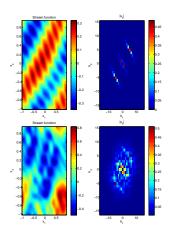
(S, Acta Numerica, 2010).  $\exists \overline{u} \in E$ :

$$I(\overline{u}) = \overline{I} := \inf\{I(u) : u \in E\}.$$

Furthermore, if  $\{u_n\}$  is a minimizing sequence satisfying  $I(u_n) \to \overline{I}$  then there is a subsequence  $\{u_{n'}\}$  that converges strongly to  $\overline{u}$  in E.



# Example: Navier-Stokes Inversion for Initial Condition



• Incompressible NSE on  $\Omega_T = \mathbb{T}^2 \times (0, \infty)$ :

$$\begin{array}{ll} \partial_t v - \nu \triangle v + v \cdot \nabla v + \nabla p = f & \text{in } \Omega_T, \\ \nabla \cdot v = 0 & \text{in } \Omega_T, \\ v|_{t=0} = u & \text{in } \mathbb{T}^2. \end{array}$$

- $y_{j,k} = v(x_j, t_k) + \eta_{j,k}, \quad \eta_{j,k} \sim N(0, \sigma^2 I_{2\times 2}).$
- $y = G(u) + \eta$ ,  $\eta \sim N(0, \sigma^2 I)$ .
- $C_0 = (-\triangle_{\text{stokes}})^{-2}$ ;  $\Phi = \frac{1}{10^3 \sigma^2} |y \mathcal{G}(u)|^2$ .



### Example: Navier-Stokes Inversion for Initial Condition

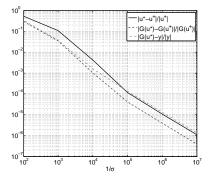


Figure: MAP estimator  $u^*$ ; Truth  $u^{\dagger}$ 



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# The Objective Functional

Recall  $\mu_0 = N(0, C_0)$  and  $\mu(du) \propto \exp(-\Phi(u))\mu_0(du)$ . Let  $\mathcal{A}$  denote a set of simple measures on H (usually Gaussian).

#### **Problem**

Find  $\nu \in \mathcal{A}$  that minimizes  $I(\nu) := D_{KL}(\nu \| \mu)$ .

Here  $D_{KL}$  =Kullbach-Leibler divergence = relative entropy

$$D_{\mathrm{KL}}(
u \| \mu) = egin{cases} \int_{\mathcal{H}} rac{d
u}{d\mu}(\mathbf{x}) \log \left(rac{d
u}{d\mu}(\mathbf{x})
ight) \mu(d\mathbf{x}) & ext{if } 
u \ll \mu \ +\infty & ext{else}. \end{cases}$$

We note, for intuition, the inequality:

$$d_{Hell}(\nu,\mu)^2 \leq 2D_{KL}(\nu\|\mu).$$



### **Existence of Minimizers**

The minimization is well-defined:

#### Theorem

(Pinski, Simpson, S, Weber, 2013) If  $\mathcal{A}$  is closed under weak convergence and there is  $\nu \in \mathcal{A}$  with  $I(\nu) < \infty$  then  $\exists \, \overline{\nu} \in \mathcal{A}$  such that

$$I(\overline{\nu}) = \overline{I} := \inf\{I(\nu) : \nu \in A\}.$$

Furthermore, if  $\{\nu_n\}$  is a minimizing sequence satisfying  $I(\nu_n) \to \overline{I}$  then there is a subsequence  $\{\nu_{n'}\}$  that converges to  $\overline{\nu}$  in the Hellinger metric:

$$d_{Hell}(\nu_n, \nu) \rightarrow 0.$$

**Example:**  $A := \mathcal{G} = \{ \text{Gaussian measures on } \mathcal{H} \}.$ 



### Parameterization of $\mathcal{G}$

Gaussian case equivalent to  $\nu = N(m, (C_0^{-1} + \Gamma)^{-1})$ .

$$J(m,\Gamma) = egin{cases} D_{\mathrm{KL}}(
u \| \mu) & \quad & \mathrm{if} \ (m,\Gamma) \in E imes \mathcal{HS}(E,E^*) \ +\infty & \quad & \mathrm{else}. \end{cases}$$

#### Theorem

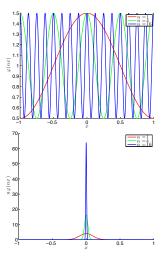
(Pinski, Simpson, S, Weber, 2013)  $\exists (\overline{m}, \overline{\Gamma}) \in E \times \mathcal{HS}(E, E^*)$ :

$$J(\overline{m},\overline{\Gamma}) = \overline{J} := \inf\{J(m,\Gamma) : (m,\Gamma) \in E \times \mathcal{HS}(E,E^*)\}.$$

Furthermore, if  $\{(m_n, \Gamma_n)\}$  is a minimizing sequence satisfying  $J(m_n, \Gamma_n) \to \overline{J}$  then there is a subsequence  $\{(m_{n'}, \Gamma_{n'})\}$  such that

$$||m_{n'} - \overline{m}||_{\mathcal{E}} + ||\Gamma_{n'} - \overline{\Gamma}||_{\mathcal{HS}(\mathcal{E}, \mathcal{E}^*)} \to 0.$$

# **Cautionary Examples**



- $\mu_0$  is Brownian bridge on [-1, 1]:
- $\nu_n := N(0, (C_0^{-1} + \Gamma_n)^{-1}).$
- Either:

$$(\Gamma_n u)(x) = \varphi(nx)u(x), \varphi \in C_{\mathrm{per}}^{\infty}, \, \mathrm{mean}\,\overline{\varphi}$$
 $\nu_n \Rightarrow \nu := N\Big(0, \big(C_0^{-1} + \overline{\varphi}Id\big)^{-1}\Big)$ 

• Or:

$$(\Gamma_n u)(x) = n\varphi(nx)u(x), \varphi \in C_0^{\infty}, \|\varphi\|_{L^1} = 1$$
  
$$\nu_n \Rightarrow \nu := N\left(0, \left(C_0^{-1} + \delta_0 Id\right)^{-1}\right).$$

# Regularization of J I

Let  $(S, \|\cdot\|_S)$  be compact in  $\mathcal{L}(E, E^*)$ .

$$J_{\delta}(\textit{m},\Gamma) = egin{cases} J(\textit{m},\Gamma) + \delta \|\Gamma\|_{\mathcal{S}}^2 & \quad & ext{if } (\textit{m},\Gamma) \in \textit{E} imes \textit{S} \ +\infty & \quad & ext{else}. \end{cases}$$

#### Theorem

(Pinski, Simpson, S, Weber, 2013)  $\exists$   $(\overline{m}, \overline{\Gamma}) \in E \times S$ :

$$J_{\delta}(\overline{m},\overline{\Gamma}) = \overline{J}_{\delta} := \inf\{J_{\delta}(m,\Gamma) : (m,\Gamma) \in E \times S\}.$$

Furthermore, if  $\{\nu_n(m_n, \Gamma_n)\}$  is a minimizing sequence satisfying  $J_\delta(m_n, \Gamma_n) \to \overline{J}_\delta$  then there is a subsequence  $\{(m_{n'}, \Gamma_{n'})\}$  such that

$$d_{Hell}(\nu_{n'}, \nu) + \|\Gamma_{n'} - \overline{\Gamma}\|_S \to 0.$$

# Regularization of J II

- Let  $H = L^2(\Omega)$ ,  $\Omega \subset \mathbb{R}^d$ ,  $C_0 = (-\triangle)^{-\alpha}$  with  $\alpha > d/2$ .
- Choose  $(\Gamma u)(x) = B(x)u(x)$  and  $S = H^r$ , r > 0.
- Thus  $\nu = N(m, C)$ ,  $C^{-1} = C_0^{-1} + B(\cdot)Id$  for potential B.

$$J_{\delta}(m,B) = egin{cases} J(m,B) + \delta \|B\|_{H^r}^2 & \qquad \text{if } (m,B) \in H^{lpha} imes H^r \ +\infty & \qquad \text{else}. \end{cases}$$

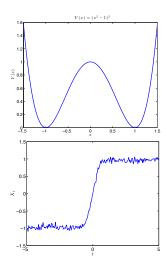
#### Theorem

(Pinski, Simpson, S, Weber, 2013)  $\exists (\overline{m}, \overline{B}) \in H^{\alpha} \times H^{r}$ :

$$J_{\delta}(\overline{m}, \overline{B}) = \overline{J}_{\delta} := \inf\{J_{\delta}(m, B) : (m, B) \in H^{\alpha} \times H^{r}\}.$$



# Example: Conditioned Diffusion in a Double Well



Consider the conditioned diffusion

$$dX_t = -\nabla V(X_t)dt + \sqrt{2\epsilon}dW_ts,.$$
  

$$X_{-T} = X_{-} < 0, X_{+T} = X_{+} > 0.$$

- For  $x_- = -x_+$ , by symmetry, we can study paths satisfying  $X_0 = 0$ ,  $X_T = x_+$
- Path space distribution approximated by N(m(t), C), with

$$C^{-1}=rac{1}{2\epsilon}\left(-rac{d^2}{dt^2}+BI
ight)$$

- B is either a constant or B = B(t)
- m and B obtained by minimizing  $D_K$



### Stochastic Root Finding & Optimization

Robbins-Monro with Derivatives and Iterate Averaging

#### Functions Estimated Via Sampling

Assume f(x) (the target function) can be estimated via F(y; x) as

$$f(x) = \mathbb{E}^{Y}[F(Y;x)], \quad f'(x) = \mathbb{E}^{Y}[\partial_{x}F(Y;x)]$$

#### Iteration Scheme (See, for instance, Asmussen & Glynn)

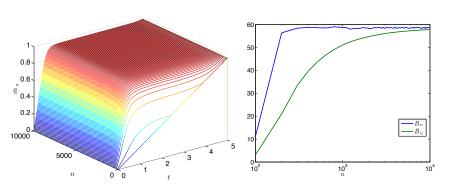
$$x_{n+1} = x_n - a_n \left(\frac{1}{M} \sum_{i=1}^{M} \partial_x F(Y_i; x_n)\right)^{-1} \left(\frac{1}{M} \sum_{i=1}^{M} F(Y_i; x_n)\right),$$

with  $a_n \sim n^{-\gamma}$ ,  $\gamma \in (\frac{1}{2}, 1)$  and  $Y_i$  i.i.d. Also let  $\bar{x}_n \equiv \sum_{j=1}^n x_j$ . Then  $\bar{x}_n \to x_*$ , with  $f(x_*) = 0$ , in distribution at rate  $n^{-1/2}$ .

### Numerical Results With Constant Potential B

 $T = 5, \epsilon = .15,$ 

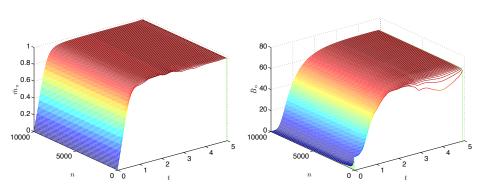
10<sup>4</sup> Iterations, 10<sup>3</sup> Samples per Iteration, 10<sup>2</sup> Points in (0, T) per Sample





### Numerical Results With Variable Potential B

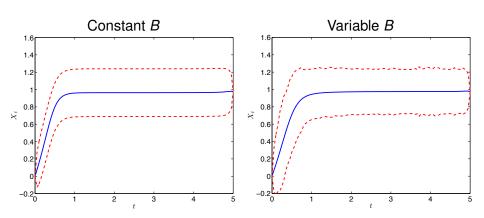
T=5,  $\epsilon=.15$ ,  $\delta=\frac{1}{2}\times 10^{-4}$ , r=1 so  $H^1$  regularization  $10^4$  Iterations,  $10^4$  Samples per Iteration,  $10^2$  Points in (0,T) per Sample





# **Model Comparison**

95% Confidence Intervals about the Mean Path





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### **MCMC**

• MCMC: create an ergodic Markov chain  $u^{(k)}$  which is invariant for approximate target  $\mu$  (or  $\mu^N$  the approximation on  $\mathbb{R}^N$ ) so that

$$\frac{1}{K}\sum_{k=1}^K f(u^{(k)}) \to \mathbb{E}^\mu f$$

- Recall  $\mu_0 = N(0, C_0)$  and  $\mu(du) \propto \exp(-\Phi(u))\mu_0(du)$ .
- Recall the Tikhonov functional  $I(u) = \frac{1}{2} ||C_0^{-\frac{1}{2}} u||^2 + \Phi(u)$ .

# Standard Random Walk Algorithm

Metropolis, Rosenbluth, Teller and Teller, J. Chem. Phys. 1953.

- Set k = 0 and Pick  $u^{(0)}$ .
- Propose  $v^{(k)} = u^{(k)} + \beta \xi^{(k)}, \quad \xi^{(k)} \sim N(0, C_0).$
- Set  $u^{(k+1)} = v^{(k)}$  with proability  $a(u^{(k)}, v^{(k)})$ .
- Set  $u^{(k+1)} = u^{(k)}$  otherwise.
- $k \rightarrow k + 1$ .

Here 
$$a(u, v) = \min\{1, \exp(I(u) - I(v))\}.$$



### New Random Walk Algorithm

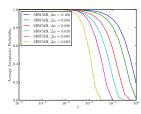
Cotter, Roberts, S and White, Stat. Sci. 2013.

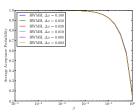
- Set k = 0 and Pick  $u^{(0)}$
- Propose  $v^{(k)} = \sqrt{(1-\beta^2)}u^{(k)} + \beta \xi^{(k)}, \quad \xi^{(k)} \sim N(0, C_0).$
- Set  $u^{(k+1)} = v^{(k)}$  with proability  $a(u^{(k)}, v^{(k)})$ .
- Set  $u^{(k+1)} = u^{(k)}$  otherwise.
- $k \rightarrow k + 1$ .

Here  $a(u, v) = \min\{1, \exp(\Phi(u) - \Phi(v))\}.$ 



# Example: Navier-Stokes Inversion for Forcing





• Incompressible NSE on  $\Omega_T = \mathbb{T}^2 \times (0, \infty)$ :

$$\begin{split} \partial_t v - \nu \triangle v + v \cdot \nabla v + \nabla p &= u & \text{in } \Omega_T, \\ \nabla \cdot v &= 0 & \text{in } \Omega_T, \\ v|_{t=0} &= v_0 & \text{in } \mathbb{T}^2. \end{split}$$

- $y_{j,k} = v(x_j, t_k) + \xi_{j,k}, \quad \xi_{j,k} \sim N(0, \sigma^2 I_{2\times 2}).$
- $y = \mathcal{G}(u) + \xi$ ,  $\xi \sim N(0, \sigma^2 I)$ .
- Prior OU process;  $\Phi = \frac{1}{\sigma^2} |y \mathcal{G}(u)|^2$ .



# Spectral Gaps

#### Theorem

(Hairer, S, Vollmer, arXiv 2012.)

- For the standard Random walk algorithm the spectral gap is bounded above by  $C N^{-\frac{1}{2}}$ .
- For the new Random walk algorithm the spectral gap is bounded below independently of dimension.



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#### What We Have Shown

#### We have shown that:

- Common Structure: A range of problems require extracting information from a probability measure on a Hilbert space, having density with respect to a Gaussian.
- Algorithmic Approaches We have laid the foundations of a range of computational methods related to this task.
- MAP Estimators Maximum a posteriori estimators can be defined on Hilbert space; there is a link to Tikhonov regularization.
- Kullback-Leibler Approximation Kullback-Leibler approximation can be defined on Hilbert space and finding the closest Gaussian results in a well-defined problem in the calculus of variations.
- Sampling MCMC methods can be defined on Hilbert space. Results in new algorithms robust to discretization.

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